

1           **ADVERSE PERINATAL CONDITIONS INCREASE THE RISK OF USING**  
2           **DISABILITY PENSION EARLY IN LIFE**

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33  
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44 **ABSTRACT**

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**Objective:** The number of young adults on disability pension is increasing in European raising questions on the related risk factors. This study aims to investigate whether adverse perinatal conditions are associated with receiving a disability pension early in life.

**Methods:** The study consisted of 453,223 individuals born in Sweden in 1973–1977, observed at ages 16–37 from 1991 through 2010. Statistics Sweden provided linked data on the children and their parents. We used logistic regression to assess the association between perinatal health conditions and receiving a disability pension, adjusting for maternal education and the sex of the child.

**Results:** New recipients of disability pension were significantly more likely to have a birth defect (Adjusted Odds Ratio [AOR] 6.63, 95% CI: 5.98 -7.34), and be small for gestational age (AOR 2.24, 95% CI: 2.17–2.85). Apgar score was significantly associated with starting to receive a disability pension at ages 16 through 18 and 19 through 29, but not at ages 30 to 33. Women had lower odds of receiving a disability pension at ages 16 to 18, however, this reversed from age 19 and upwards. Persons with higher maternal education were less likely to receive a disability pension compared to persons with  $\leq 9$  years of maternal education level. Overall, the effects of the studied perinatal health conditions were strongest in those 16 to 18 years at disability pension, but reduced as age increased.

**Conclusion:** Having a birth defect was the strongest indicator of receiving a disability pension during early adulthood, followed by small for gestational age, and low Apgar score. Our findings suggest that policies and programs geared at promoting optimal health at birth might improve overall health over the lifespan, contributing to a reduction in receiving early disability pensions, and dependence on health services and social welfare.

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91   **INTRODUCTION**

92   Over the past few decades, Europe has witnessed an increase in the number of people  
93   receiving disability pensions, [1, 2] with several countries reporting an increase in the number  
94   of young adults as new recipients. [2] Disability pensions are a social security scheme that  
95   provides income support to people of working age with long-term limitations in their working  
96   capacity due to ill health. [3, 4] Disability pension is an important part of the public support  
97   programs for people with disabilities in Sweden. [3, 4] Receiving a disability pension has also  
98   been viewed as an indicator of long-term ill health, and once on disability pension, very few  
99   recipients return to active work. [1] This trend of early exit from the labor force via disability  
100   pension is highly unsustainable over time, as it increases financial pressures on the  
101   government, and aggravates the anticipated future labor force shortage due to the aging  
102   population. [2]

103   Several studies have attempted to understand the factors associated with the risk of utilizing a  
104   disability pension. A growing body of research has identified several socio-economic and  
105   health factors associated with the receipt of disability pensions. [2, 5] Some of these identified  
106   adulthood socio-demographic risk factors included education, occupation, civil status, family  
107   structure, and place of residence.[6-12] Individuals receiving disability pensions also have  
108   more adverse health outcomes, such as poor self-rated health, alcohol use, frequent use of  
109   primary health care, and noted genetic differences. [7, 10, 13-15]

110   A few studies using the life course critical model [16] have also investigated the link between  
111   childhood conditions and receiving a disability pension later in life. The critical model

112 suggests that suboptimal perinatal conditions cause long-lasting changes in the developing  
113 organ structures, and in the functioning of biological systems, which in turn places an  
114 individual at an increased risk of chronic diseases during adulthood. [16] Additionally, a  
115 handful of studies established a link between childhood socio-economic position and the risk  
116 of having a disability later in life. [17, 18] Some studies noted that receiving a disability  
117 pension during adulthood was higher among persons with low birth weight, [19, 20] and  
118 among those born small for gestational age. [11] However, this evidence on the linkage  
119 between perinatal health and the receipt of a disability pension during early adulthood is still  
120 insufficient.

121 From our literature search, we identified no single study that has investigated the association  
122 between receiving a disability pension and having a birth defect, and a low Apgar score. We  
123 think that these associations are worthy of further investigation, as some evidence suggests  
124 that persons with birth defects are more likely to report a developmental disability later in life.  
125 [21-23] Some studies have reported a link between a low Apgar score at five-minutes and  
126 minor disabilities at school age. [24] A low Apgar score was also associated with a neurologic  
127 disability in early adulthood. [25, 26] Factors associated with an increased risk of any form of  
128 disability affect one's overall quality of health, and as such, are likely to increase the risk of a  
129 work disability that leads to receiving a disability pension. Thus, we hypothesize that having a  
130 birth defect and a low Apgar score is associated with receiving a disability pension. To test  
131 this hypothesis, we followed the birth cohort of 1973–1977 from ages 16 to 37, with the aim  
132 of investigating the association between perinatal health factors, as measured by birth defects,  
133 Apgar score, and being small for gestational age, and the receipt of a Swedish disability  
134 pension during early adulthood.

135

136 **MATERIALS AND METHODS**

137 The study population consisted of 453,223 individuals belonging to five complete birth-year  
138 cohorts from 1973 to 1977, who were in Sweden between ages 16–37, during the study period  
139 of 1991 to 2010. Initially, we had identified 693,247 individuals belonging to this birth  
140 cohort, but 240,024 were excluded, as some were born outside of Sweden, some died, and  
141 others emigrated either before or during the study period (details in Figure 1).

142 (Figure 1 here)

143 The Swedish Medical Birth Register provided information on the total newborn population,  
144 and their perinatal health outcomes. This register also collects information on all congenital  
145 anomalies observed during the first year of life. [27] We obtained information on sex,  
146 mother’s education level, and receipt of disability benefits from the Longitudinal Integration  
147 Database for Health Insurance and Labour Market Studies (LISA database). Statistics Sweden  
148 linked the index person’s data and the data of their mother obtained from these two data  
149 sources, using the unique Swedish personal identity number (PIN)- After data linkage,  
150 Statistics Sweden made the data anonymous before delivering it to the Swedish Initiative for  
151 Research on Microdata in Social and Medical Sciences (Umeå SIMSAM Lab), [28] where we  
152 performed all analyses.

153 ***Study variables***

154 *Receiving a disability pension:* This outcome variable was measured when the individuals  
155 were between the ages of 16 to 37. Sweden uses a systematic medical examination, as  
156 codified by Swedish social security legislation, to measure diminished health and work  
157 capacity when assessing eligibility to receive this financial benefit. [3-5] However, Swedish  
158 disability pension legislation has frequently changed. From 1991–2002, individuals were

159 eligible to receive a disability pension if they were between ages 16–64, with medical  
160 evidence confirming their inability to work due to chronic ill health. [3, 4] From 2003  
161 onwards, the basis for granting this financial security remained the same, but the minimum  
162 age limit was raised to age 19, and the term disability pension was replaced with activity  
163 compensation, which is payable to persons aged 19–29, and sickness compensation, payable  
164 to persons aged 30–64. In this paper, we use the term disability pension as an umbrella term  
165 that includes disability pension, activity compensation, and sickness compensation. We  
166 recorded individuals as having received a disability pension from their first year of receiving  
167 the benefit, “yes” for those who received a disability pension and “no” for those who did not.

168 When selecting explanatory variables, we included perinatal health indicators because  
169 previous research suggests a link, [19–23] and/or because we considered them plausible risk  
170 factors for receiving a disability pension. We obtained the variable birth defect from the  
171 Medical Birth Register, already coded in the register as yes and no. The variable Apgar score  
172 at five minutes measures the physical condition of the newborn at 5 minutes after birth on a  
173 scale of 0–10. [29] We categorized this variable according to existing criteria, which  
174 considers a total score of less than 7 as a low Apgar score, while a score within a range of 7 to  
175 10 is considered normal. [29, 30] The variable gestational age was pre-categorized into two  
176 groups, based on the Swedish growth standards that account for both birth weight for  
177 gestational age and sex. The 5<sup>th</sup> percentile (z-score below -1.64) was the threshold, individuals  
178 below this were pre-categorized as small for gestational age, and those above considered  
179 appropriate for gestational age. We categorize sex as either man or woman, and maternal  
180 education is categorized into  $\leq 9$  years of school (reference category), upper secondary  
181 education, and university education.

182

183 *Statistical analyses*

184 In the descriptive analysis, we used cross tabulation to compare the explanatory variables  
185 between individuals who received a disability pension, and those who did not (Table 1). We  
186 examined bivariable correlation between birth defect and all other independent variables, and  
187 found no evidence of a strong correlation; thus, we chose to keep all of the independent  
188 variables. We conducted logistic regression to assess whether any of the perinatal conditions  
189 were associated with the odds of receiving a disability pension between the ages of 16–37.  
190 We examined the independent association between each of the variables, and the odds of  
191 receiving a disability pension (see Bivariable results column in Table 2). In Table 2, Models  
192 1–3, we assessed changes in the odds of receiving a disability pension by adjusting for  
193 covariates. In Model 1, we included all perinatal health variables; in Model 2, we added sex  
194 and maternal education to Model 1, and in Model 3, we checked for the interactions between  
195 sex and all other explanatory variables.

196 Furthermore, we performed logistic regression analyses to account for the different ages  
197 beginning to receive a disability pension (See Table 3). We created separate models using an  
198 individual's age at the start of their receiving a disability pension, taking into consideration  
199 the changes in the national legislation on disability pensions, i.e. ages 16–18, 19–29, and 30–  
200 33. Model 4 measured the odds of receiving a disability pension at ages 16–18, Model 5  
201 estimated the odds of receiving a disability pension at ages 19–29, and Model 6 assessed the  
202 odds of receiving a disability pension at ages 30–33. In further analyses, (See Table 4,  
203 appendix) we introduced all explanatory variables and interaction terms simultaneously for  
204 each of the above that was presented in Table 3. We assessed multi-collinearity for all  
205 adjusted models by calculating the variance inflation factor (VIF), regressing each  
206 independent variable on all the other independent variables, and found no strong evidence of

207 multi-collinearity. Using Anova, we evaluated the overall fitness of the model. Odds ratios  
 208 (ORs) with 95% confidence intervals (CIs) are reported. Statistical significance was attained  
 209 with a  $p < 0.05$ . We performed our statistical analyses using the R software.

## 210 RESULTS

211 We present the differences in the study population's characteristics in Table 1. The total  
 212 number of people that began receiving a disability pension between ages 16-37 years, during  
 213 the duration of the follow-up, was 18,854 (4% of the 453,223 participants). The proportion of  
 214 individuals with birth defects who received a disability pension was twice as large when  
 215 compared to individuals without birth defects (8% vs. 4%, respectively). Disability pension  
 216 reception was more common among females than males, 5% versus 3%, respectively. The  
 217 prevalence of disability pension reception was highest among those with maternal education  
 218 less  $\leq 9$  years of schooling (5%), however, the prevalence was similar among those with  
 219 mothers with upper secondary or university education (3%).

220 **Table 1. Perinatal characteristics of the birth cohort of 1973–1977 by disability pension status (n=453,223)**

<i>Description</i>	<i>No Disability Pension</i> <i>N=434,369 (96%)</i>		<i>Disability Pension</i> <i>N=18,854 (4%)</i>	
		<i>(%)</i>		<i>(%)</i>
<b><i>Birth defect</i></b>				
<i>No</i>	411,498	(96)	16,939	(4)
<i>Yes</i>	19,523	(92)	1,739	(8)
<i>Data Missing</i>	3,348	(95)	176	(5)
<b><i>Apgar at 5 minutes</i></b>				
$\geq 7$	311,306	(96)	13,159	(4)
$< 7$	4,406	(91)	428	(9)
<i>Data Missing</i>	118,657	(96)	5,267	(4)
<b><i>Small gestational age</i></b>				
<i>No</i>	403,644	(96)	16,767	(4)
<i>Yes</i>	18,516	(93)	1,422	(7)
<i>Data Missing</i>	12,209	(95)	665	(5)
<b><i>Sex</i></b>				
<i>Male</i>	226,431	(97)	7,941	(3)
<i>Female</i>	207,938	(95)	10,913	(5)
<b><i>Mother's education</i></b>				
$< 9$ years	178,468	(95)	8,675	(5)
<i>Upper secondary</i>	159,993	(97)	5,628	(3)
<i>University</i>	32,717	(97)	966	(3)
<i>Data Missing</i>	63,191	(95)	3,585	(5)



221 In Table 2, all of the bivariable results showed significant associations. Individuals with a  
222 birth defect were at higher odds of receiving a disability pension, compared with their  
223 counterparts without a birth defect (OR 2.16, 95% CI: 2.05–2.28). Those with low Apgar  
224 scores were more likely to receive disability pensions compared to those with Apgar scores of  
225 7–10 (OR 2.29, 95% CI: 2.08–2.54). Persons who were small for gestational age were more  
226 likely to receive a disability pension compared to those that were not. Females had higher  
227 odds of receiving a disability pension compared to males (OR 1.85, 95% CI: 1.75–1.95).  
228 Using  $\leq 9$  years of maternal schooling as the reference category, those born to mothers with a  
229 high level of education were less likely to use the disability pension.

230 In Model 1, simultaneous adjusting for the three perinatal health variables confirmed  
231 significantly increased odds of receiving a disability pension among persons with birth  
232 defects, a low Apgar score and who were small for gestational age. In Model 2, we added sex  
233 and maternal education to the previous model, and noted that birth defects, low Apgar score,  
234 being small for gestational age, and being a woman remained significantly associated with  
235 higher odds of receiving a disability pension, while high maternal education was significantly  
236 associated with lower odds of receiving a disability pension. In Model 3, we added interaction  
237 terms to the previous model (Model 2), the main effect remained significant and in similar  
238 direction as observed in the previous model. Model 3 also revealed an interaction between  
239 birth effect and sex and an interaction between birth effect and small for gestation age. Using  
240 Anova, we evaluated the overall fitness of the models, and found significant evidence of the  
241 overall effects of all independent variables on the dependent variable.

242

243

245 **Table 2. The unadjusted and adjusted associations between perinatal factors and the receipt of a disability**  
 246 **pension**

Perinatal factors	Bivariable results		Multivariable results	
	OR (95% CI)	Model 1: (n=322,464) OR (95% CI)	Model 2: (n=273,708) OR (95% CI)	Model 3: (n=273,708) OR (95% CI)
<b>Birth defect</b>				
No	1.00	1.00	1.00	1.00
Yes	2.16 (2.05–2.28) ***	2.31 (2.18–2.45) ***	2.52 (2.36–2.69) ***	2.74 (2.49–3.00)***
<b>Apgar score</b>				
7-10	1.00	1.00	1.00	1.00
<7	2.29 (2.08–2.54) ***	2.12 (1.91–2.35)***	2.19 (1.95–2.47)***	2.12 (1.77–2.52)***
<b>Small for gestational age</b>				
No	1.00	1.00	1.00	1.00
Yes	1.85 (1.75–1.95)***	1.84 (1.72–1.96) ***	1.73 (1.61–1.86) ***	1.73 (1.54–1.94)***
<b>Sex</b>				
Male	1.00		1.00	1.00
Female	1.49 (1.45–1.54)***		1.54 (1.48–1.59)***	1.55 (1.46–1.64)***
<b>Mother's education</b>				
<9 years	1.00		1.00	1.00
Upper secondary	0.72 (0.69–0.75) ***		0.76 (0.73–0.79)***	0.74 (0.69–0.79)***
University	0.61 (0.57–0.65)***		0.63 (0.58–0.68)***	0.67 (0.59–0.75)***
<b>Birth defect*sex</b>				
Male: No birth defect				1.00
Female: birth defect				0.85 (0.78–1.02)
<b>Sex*Apgar score</b>				
Male: Apgar 7-10				1.00
Female: Apgar <7				1.07 (0.84–1.35)
<b>Sex*Gestational age</b>				
Male: not SGA				1.00
Female: SGA				1.00 (0.86–1.16)
<b>Mother's education*sex</b>				
Male: <9 years				1.00
Female: Upper secondary				1.04 (0.96–1.13)
Female: University				0.90 (0.77–1.06)
<b>Birth defect*Apgar score</b>				
Birth defect No*Apgar score 7-10				1.00
Birth defect Yes*Apgar score <7				1.26 (0.89–1.76)
<b>Birth defect *SGA</b>				
Birth defect No*SGA No				1.00
Birth defect Yes*SGA Yes				1.39 (1.12–1.74)*
<b>SGA*Apgar score</b>				
SGA No*Apgar score 7-10				1.00
SG Yes*Apgar score <7				0.86 (0.63–1.16)

247 *Model 1 contains perinatal health variables, Model 2 adds sex and mother's education level to model 1, Model 3*  
 248 *extends Model 2 by including interaction effects. \*\*\* indicates p-value <0.001, \* p-value<0.05*

249 In Table 3, we present our investigation of the association between perinatal conditions and  
 250 the odds of receiving a disability pension in three different age groups. Among those who  
 251 started receiving a disability pension between ages 16–18 (Model 4), the new recipients were  
 252 significantly more likely to have a birth defect, a low Apgar score, and to be small for  
 253 gestational age. Model 5 shows the increased odds of receiving a disability pension between

254 ages 19–29 among those with a birth defect, a low Apgar score, who were small for  
 255 gestational age, and women. In Model 6, starting to receive a disability pension between ages  
 256 30–33 was associated with having a birth defect and being small for gestational age, but not  
 257 with a low Apgar score. High maternal education was associated with lower odds of receiving  
 258 a disability pension for all age groups presented in Table 3.

259 **Table 3 shows the associations between perinatal factors and disability pension reception in**  
 260 **the 1973–1977 birth cohort, stratified by age at the start of receiving disability pension and**  
 261 **adding all possible interactions**

	Model 4: Age 16- 18 years OR (95% CI) (n=)	Model 5: 19- 29 years OR (95% CI) (n=269,716)	Model 6: 30-33 years OR (95% CI) (n=)
<b>Birth defect</b>			
No	1.00	1.00	1.00
Yes	5.89 (5.06–6.84)***	1.49 (1.24–1.79)***	1.39 (1.08–1.78)**
<b>Apgar score</b>			
7-10	1.00	1.00	1.00
<7	4.25 (3.12–5.66)***	1.53 (1.41–1.67)*	1.25 (0.75–1.97)
<b>Small for gestational age</b>			
No	1.00	1.00	1.00
Yes	2.15 (1.69–2.69)**	1.49 (1.22–1.80)***	1.43 (1.09–1.84)***
<b>Sex</b>			
Male	1.00	1.00	1.00
Female	0.73 (0.64–0.85)***	1.53 (1.41–1.67)***	2.16 (1.95–2.40) ***
<b>Mother's education</b>			
<9 years	1.00	1.00	1.00
Upper secondary	0.75 (0.66–0.85)***	0.71 (0.65–0.79)***	0.72 (0.64–0.82)***
University	0.76 (0.59–0.96)*	0.66 (0.54–0.79)***	0.56 (0.43–0.73)***
<b>Interaction terms</b>			
<b>Birth defect*sex</b>			
Male: No birth defect	1.00	1.00	1.00
Female: birth defect	1.27 (1.03–1.58)*	0.99 (0.78–1.26)	0.88 (0.64–1.21)
<b>Apgar score*sex</b>			
Male: Apgar 7-10	1.0	1.00	1.0
Female: Apgar <7	1.22 (0.82–1.80)	1.46 (0.98–2.20)	0.85 (0.47–1.55)
<b>Sex*Gestational age</b>			
Male: not SGA	1.00	1.00	1.00
Female: SGA	1.39 (1.04–1.86)**	1.02 (0.81–1.30)	0.98 (0.72–1.35)
<b>Mother's education*sex</b>			
Male: <9 years	1.00	1.00	1.00
Female: Upper secondary	1.21 (0.99–1.47)	1.11 (0.98–1.25)	0.97 (0.83–1.14)
Female: University	1.25 (0.88-1.77)	0.93 (0.73–1.19)	0.84 (0.59–1.18)
<b>Birth defect*Apgar score</b>			
Birth defect No*Apgar score 7-10	1.00	1.00	1.00
Birth defect Yes*Apgar score <7	0.90 (0.57–1.40)	0.86 (0.39–1.67)	1.07 (0.31–2.76)
<b>Birth defect *SGA</b>			
Birth defect No*SGA No	1.00	1.00	1.00
Birth defect Yes*SGA Yes	1.15 (0.83–1.58)	1.27 (0.82–1.89)	1.06 (0.55–1.86)
<b>SGA*Apgar score</b>			
SGA No*Apgar score 7-10	1.00	1.00	1.00
SG Yes*Apgar score <7	0.56 (0.34–0.91)*	0.84 (0.49–1.37)	1.29 (0.61–2.52)

262 Models 4–6 contain all the studied perinatal variables, showing the main effect and interaction effects.

263 \*\*\* indicates p-value <0.001, \* p-value<0.01 and \* p-value<0.05

264

265 **DISCUSSION**

266 Our findings support the study's main hypothesis that having a birth defect is significantly  
267 associated with beginning to receive a disability pension during early adulthood. A low Apgar  
268 score was associated with receiving a disability pension before age 30, but not afterwards. We  
269 confirmed an association that was observed earlier between being small for gestational age  
270 and increased odds of receiving a disability pension. Women were less likely to receive a  
271 disability pension between ages 16–18, but had increased odds of receiving a disability  
272 pension from age 19 onwards. We further observed that the effect of perinatal health was  
273 strongest among those who started to receive a disability pension between ages 16–18, but the  
274 strength of the association reduced as age when beginning to receive a disability pension  
275 increased, even though this effect remained statistically significant. Compared to persons with  
276 maternal education <9 years, individuals with higher maternal education level were  
277 significantly less likely to receive a disability pension as age increased.

278 The strength of our study stems mainly from using register data with a nationwide coverage,  
279 which ensured high completeness, limited follow-up loss, and no recall bias. This study  
280 suffers from some limitations, such as a potential selection bias relating to death as an  
281 exclusion criterion. Children with adverse perinatal health outcomes are most likely to die,  
282 and hence the exclusion of children who died might have led to underestimations of the effect,  
283 and potentially a selection bias. However, we still observed high odds among those with  
284 adverse health indicators. Our results reflect the situation among those who were alive from  
285 ages 16 to 37. The change in the minimum age eligibility for the disability pension, from age  
286 16 to age 19 in 2002, might have led to classification bias in our study. However, we consider  
287 this a minor problem, because the eligibility criteria remained based on the presence of a  
288 long-term work disabling health condition, both prior to and beyond 2002. [39] Additionally,

289 we considered the legislation change by analyzing data for the different age groups separately,  
290 i.e. ages 16–18, 19–29, and 30–33.

291 As far as we know, ours is the first study to investigate the associations between birth defects,  
292 Apgar score at 5 minutes, and receiving a disability pension between ages 16–37. Our  
293 findings confirm our hypothesis that having a birth defect and low Apgar score is associated  
294 with increased odds of receiving a disability pension early in life. These observations are  
295 biologically plausible, as existing literature already suggests a link between birth defects, [21,  
296 23] low Apgar score at 5 minutes, and disability indicators such as a developmental disability.  
297 [25, 26] Our finding that individuals who were small for gestational age were more likely to  
298 receive a disability pension during early adulthood supports earlier studies that report similar  
299 associations. [11, 19, 20] We observed interaction between birth defect and being born small  
300 for gestational. The increased odds of receiving disability pension for individuals with  
301 adverse perinatal health might imply ill health over the lifespan, given that prolonged ill  
302 health is a legal requirement for receiving a disability pension. [5, 31]

303 We report that the effects of the perinatal health factors appeared to weaken as age at  
304 beginning to receive a disability pension increased. This might imply that individuals who had  
305 severe health problems needed to start receiving a disability pension earlier. It could also be  
306 that, as these individuals get older, the effects of adulthood exposures become pronounced,  
307 outweighing the effects of the perinatal health factors. Future studies extending the model to  
308 include adult factors could help clarify this association.

309 The fact that higher odds of receiving a disability pension were associated with low maternal  
310 education concurs with the theory of fetal origin. Low education possibly indicates a poor  
311 maternal socio-economic position, predisposing offspring to further socio-economic,  
312 behavioral, and pathological disadvantages across the lifespan [16, 19] that eventually lead to

313 receipt of a disability pension. Our finding that women were more likely to receive a  
314 disability pension is in line with reports from earlier studies in Sweden [1, 2], and from other  
315 European countries. [31] We observed that women had lower odds of receiving a disability  
316 pension at ages 16–18, but had higher odds from age 19 and onwards. Earlier studies have  
317 also reported gender differences in the use of disability pensions. [1, 6, 20, 31, 32] Gravseth  
318 noted an increase in women’s incidence rates beginning in their late twenties and onwards,  
319 but not before this age. [20] The explanation for these gender differences is not clear, but  
320 previous research has attributed it to the underlying gender structure that dictates gendered  
321 living and working conditions, exposing women engaging in paid work to a “double burden”  
322 that results from combining work and responsibility for the home and children. As a result,  
323 women’s health tends to suffer as they reach the age that requires combining gainful  
324 employment with family life. [33, 34] However, these observed differences could perhaps be  
325 due to other health and socio-economic conditions outside the scope of this study.

326

## 327 **CONCLUSION**

328 This study provides evidence of an association between birth defects, low Apgar score at 5  
329 minutes, and the increased odds of receiving a disability pension, taking other early life health  
330 and maternal measures into consideration. The confirmed association points to the complexity  
331 in the relationship between early life conditions and the later receipt of a disability pension.  
332 Our findings contribute to previous knowledge on the predictors of disability pension receipt,  
333 highlighting a need for better-focused strategies to promote early health, as this could  
334 contribute to reduced work incapacity during the early stages of adulthood. This is critical to  
335 consider, because evidence suggests that the majority of people that start to receive a  
336 disability pension tend to do so on a long-term basis. Among females, increase in age was

337 associated with a higher use of a disability pension. Our findings suggest that those with poor  
338 perinatal outcomes are likely to be more susceptible to disabling chronic conditions in early  
339 adulthood, reducing their work capacity and hence fostering a need to receive a disability  
340 pension. The fact that those with poor health indicators at birth, and children of mothers with  
341 low education, were more likely to receive a disability pension suggests a need for a  
342 continued review of public and social policies aimed at improving early life conditions. This  
343 would contribute to a reduction in the number of disability pension recipients, and to an  
344 improvement in overall societal health and well-being.

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